

THE FOURIER TRANSFORM OF THE MULTIDIMENSIONAL  
GENERALIZED GAUSSIAN DISTRIBUTION

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**Abstract:** We present expressions for the generalized Gaussian distribution in  $n$  dimensions and compute their Fourier transforms. We obtain expressions in terms of Bessel functions and Maclaurin series.

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**Key Words:** generalized Gaussian distribution, Fourier transform, characteristic function, Bessel function, Maclaurin series

1. Introduction

The generalized Gaussian (GG) or the generalized normal (GN) distribution [38], also known as the exponential power distribution [6] or the generalized Laplace distribution [18, 20, 22], was first proposed by Subbotin [31] as a generalized error distribution. Statistical problems have been considered for this distribution by many authors. For example, parameter estimation was studied in [2, 9, 17, 25, 34, 36, 37, 39], Bayesian statistical analysis in [4, 5, 6, 7, 10, 35], and other statistical problems in [8, 14, 15, 27, 28, 29, 30, 32, 33]. The GG distribution is also used in several applications related to signal and image analysis, for example [3, 11, 24, 26]. For other information concerning this distribution and its applications see [6, 18, 20, 22].

The Fourier transform (or the characteristic function) of the GG distribution is very often used in the applications. In one dimension (1D), the GG distribution is defined for  $x \in \mathbb{R}$  by the formula

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$$f_\lambda(x) = e^{-|x|^\lambda} \quad (1)$$

where  $\lambda$  is a strictly positive real number. Its Fourier transform is defined by

$$\hat{f}_\lambda(\omega) = \int_{-\infty}^{+\infty} e^{-|x|^\lambda} e^{-2\pi i \omega x} dx. \quad (2)$$

Series expansions have been obtained in this case for integer valued parameter (see [12]) and more generally for real valued parameter, see [13].

In this paper we propose some generalizations in  $n$  dimensions ( $nD$ ) of the GG distribution and study their Fourier transforms. We extend some results presented in [15, 23]. Let  $x^t = (x_1, x_2, \dots, x_n)$  and  $\omega^t = (\omega_1, \omega_2, \dots, \omega_n)$  be elements of  $\mathbb{R}^n$ , and let  $\omega^t x$  be the standard scalar product in  $\mathbb{R}^n$  defined by

$$\omega^t x = \sum_{i=1}^n \omega_i x_i.$$

If  $f(x) : \mathbb{R}^n \rightarrow \mathbb{R}$ , its Fourier transform,  $\hat{f}_\lambda(\omega)$ , is given by

$$\hat{f}_\lambda(\omega) = \int_{\mathbb{R}^n} f_\lambda(x) e^{-2\pi i \omega^t x} dx, \quad (3)$$

where  $dx$  stands for  $dx_1 dx_2 \cdots dx_n$ .

## 2. First Generalization

We consider the simplest generalization, the separable case with  $n$  possible different parameters. We set

$$F_{\lambda_1, \lambda_2, \dots, \lambda_n}(x_1, x_2, \dots, x_n) = f_{\lambda_1}(x_1) f_{\lambda_2}(x_2) \dots f_{\lambda_n}(x_n). \quad (4)$$

This form is not invariant by rotation but is separable, as a consequence its Fourier transform is the product of  $n$  1D Fourier transforms. Indeed we have

$$\hat{F}_{\lambda_1, \lambda_2, \dots, \lambda_n}(\omega_1, \omega_2, \dots, \omega_n) = \hat{f}_{\lambda_1}(\omega_1) \hat{f}_{\lambda_2}(\omega_2) \dots \hat{f}_{\lambda_n}(\omega_n), \quad (5)$$

where formulas for  $\hat{f}_\lambda(\omega)$  will be given in the next section.

### 3. Second Generalization

#### 3.1. Definition

The second and natural generalization of the GG distribution is defined for  $x \in \mathbb{R}^n$  by the formula

$$G_\lambda(x) = e^{-\|x\|^\lambda} = e^{-(x_1^2+x_2^2+\dots+x_n^2)^{\lambda/2}}, \tag{6}$$

where  $\|x\|$  is the standard norm in  $\mathbb{R}^n$  defined by

$$\|x\| = (x^t x)^{\frac{1}{2}} = \left( \sum_{i=1}^n x_i^2 \right)^{\frac{1}{2}}.$$

So defined, the function  $G_\lambda(x)$  is a radial function and is invariant by rotation.

Since  $G_\lambda(x)$  is a rapidly decaying function, its Fourier transform and all its derivatives are well defined, see [19, 21]. The Fourier transform of  $G_\lambda(\omega)$  is given by

$$\hat{G}_\lambda(\omega) = \int_{\mathbb{R}^n} e^{-\|x\|^\lambda} e^{-2\pi i \omega^t x} dx. \tag{7}$$

#### 3.2. First Formula

Let us observe that for any linear orthogonal transformation on  $\mathbb{R}^n$ ,  $V : \mathbb{R}^n \rightarrow \mathbb{R}^n$  such that  $V^t V = I$ , we have  $x = Vy$  or  $V^t x = y$ , and

$$\hat{G}_\lambda(\omega) = \int_{\mathbb{R}^n} e^{-\|y\|^\lambda} e^{-2\pi i \omega^t Vy} dy$$

since  $|\text{Det}(V)| = 1$ . For any  $\omega \neq 0$ , if we consider a linear orthogonal transformation  $V$  such that its first column is  $\frac{1}{\|\omega\|}\omega$ , then  $\omega^t Vy = \|\omega\| y_1$  because  $\omega$  is orthogonal to the  $(n - 1)$  other columns of  $V$ . Then, for any  $\omega$  we can write

$$\hat{G}_\lambda(\omega) = \int_{\mathbb{R}^n} e^{-\|y\|^\lambda} e^{-2\pi i \|\omega\| y_1} dy.$$

Let us now consider the three cases:

(i)  $n = 1$ : We have

$$\hat{G}_\lambda(\omega) = \int_{-\infty}^{+\infty} e^{-|y|^\lambda} e^{-2\pi i \|\omega\| y} dy$$

$$= 2 \int_0^{+\infty} e^{-y^\lambda} \cos(-2\pi|\omega|y) dy.$$

(ii)  $n = 2$ : We can write

$$\begin{aligned} \hat{G}_\lambda(\omega) &= \int_{\mathbb{R}^2} e^{-(y_1^2+y_2^2)^{\frac{\lambda}{2}}} e^{-2\pi i\|\omega\|y_1} dy_1 dy_2 \\ &= 2 \int_{-\infty}^{+\infty} e^{-2\pi i\|\omega\|y} \int_0^{+\infty} e^{-(y^2+\rho^2)^{\frac{\lambda}{2}}} d\rho dy, \end{aligned}$$

where we used  $y_1 = y$  and  $y_2 = \rho$ .

(iii)  $n \geq 3$ : Let us consider the change of variables

$$\begin{cases} y_1 &= y, \\ y_2 &= \rho \cos(\theta_2), \\ y_3 &= \rho \sin(\theta_2) \cos(\theta_3), \\ \vdots &\vdots \\ y_{n-1} &= \rho \sin(\theta_2) \sin(\theta_3) \cdots \sin(\theta_{n-2}) \cos(\theta_{n-1}), \\ y_n &= \rho \sin(\theta_2) \sin(\theta_3) \cdots \sin(\theta_{n-2}) \sin(\theta_{n-1}), \end{cases}$$

where  $y \in \mathbb{R}$ ,  $\rho \in [0, +\infty)$ ,  $\theta_2, \theta_3, \dots, \theta_{n-2} \in [0, \pi]$ , and  $\theta_{n-1} \in [-\pi, \pi]$ . We have

$$\begin{aligned} dy_1 dy_2 \cdots dy_n \\ = \rho^{n-2} \sin^{n-3}(\theta_2) \sin^{n-2}(\theta_3) \cdots \sin(\theta_{n-2}) dy d\rho d\theta_2 d\theta_3 \cdots d\theta_{n-2} d\theta_{n-1}, \end{aligned}$$

it follows that

$$\hat{G}_\lambda(\omega) = S_{n-1} \int_{-\infty}^{+\infty} e^{-2\pi i\|\omega\|y} \int_0^{+\infty} e^{-(y^2+\rho^2)^{\frac{\lambda}{2}}} \rho^{n-2} d\rho dy,$$

where  $S_{n-1}$  is the area of the unit sphere in  $\mathbb{R}^{n-1}$  given by

$$S_{n-1} = \int_{[0,\pi]^{n-3} \times [-\pi,\pi]} \prod_{j=2}^{n-2} \sin^{n-1-j}(\theta_j) d\theta_2 d\theta_3 \cdots d\theta_{n-2} d\theta_{n-1} = \frac{2\pi^{\frac{n-1}{2}}}{\Gamma(\frac{n-1}{2})}.$$

Hence

$$\hat{G}_\lambda(\omega) = \frac{2\pi^{\frac{n-1}{2}}}{\Gamma(\frac{n-1}{2})} \int_{-\infty}^{+\infty} e^{-2\pi i\|\omega\|y} \int_0^{+\infty} e^{-(y^2+\rho^2)^{\frac{\lambda}{2}}} \rho^{n-2} d\rho dy.$$

It follows that for  $n \geq 2$

$$\hat{G}_\lambda(\omega) = \frac{2\pi^{\frac{n-1}{2}}}{\Gamma(\frac{n-1}{2})} \int_{-\infty}^{+\infty} \cos(2\pi \|\omega\| y) \int_0^{+\infty} e^{-(y^2+\rho^2)^{\frac{\lambda}{2}}} \rho^{n-2} d\rho dy. \tag{8}$$

Using the change of variables

$$\begin{cases} y &= \varsigma \cos(\psi), \\ \rho &= \varsigma \sin(\psi), \end{cases}$$

where  $\varsigma \in [0, +\infty)$  and  $\psi \in [0, \pi]$ , we have  $d\rho dy = \varsigma d\varsigma d\psi$ , and (8) becomes

$$\hat{G}_\lambda(\omega) = \frac{2\pi^{\frac{n-1}{2}}}{\Gamma(\frac{n-1}{2})} \int_0^{+\infty} e^{-\varsigma^\lambda} \varsigma^{n-1} \int_0^\pi \sin^{n-2}(\psi) \cos(2\pi \|\omega\| \varsigma \cos(\psi)) d\psi d\varsigma.$$

From the series expansion of the  $\cos(\cdot)$ , we have

$$\begin{aligned} \int_0^\pi \sin^{n-2}(\psi) \cos(2\pi \|\omega\| \varsigma \cos(\psi)) d\psi \\ = \sum_{k=0}^{+\infty} (-1)^k \frac{(2\pi \|\omega\| \varsigma)^{2k}}{(2k)!} \int_0^\pi \sin^{n-2}(\psi) \cos^{2k}(\psi) d\psi. \end{aligned}$$

Then we use the following identities (see [1])

$$\int_0^\pi \sin^{n-2}(\psi) \cos^{2k}(\psi) d\psi = \frac{\Gamma(\frac{2k+1}{2})\Gamma(\frac{n-1}{2})}{\Gamma(\frac{2k+n}{2})},$$

and

$$\Gamma\left(\frac{2k+1}{2}\right) = \frac{(2k)!}{2^{2k} k!} \pi^{\frac{1}{2}},$$

to obtain

$$\hat{G}_\lambda(\omega) = 2\pi^{\frac{n}{2}} \int_0^{+\infty} e^{-\varsigma^\lambda} \varsigma^{n-1} \left[ \sum_{k=0}^{+\infty} (-1)^k \frac{(\pi \|\omega\| \varsigma)^{2k}}{\Gamma(k+1)\Gamma(\frac{2k+n}{2})} \right] d\varsigma, \tag{9}$$

which is valid for any  $n \geq 1$ .

### 3.3. Fourier Transform and Bessel Function

Introducing the Bessel function

$$J_\nu(\xi) = \left(\frac{\xi}{2}\right) \sum_{k=0}^{+\infty} (-1)^k \frac{\xi^{2k}}{2^{2k} k! \Gamma(\nu + k + 1)},$$

with  $\nu = \frac{n}{2} - 1$  and  $\xi = 2\pi \|\omega\| \varsigma$ , (9) becomes

$$\hat{G}_\lambda(\omega) = \frac{2\pi}{\|\omega\|^{\frac{n}{2}-1}} \int_0^{+\infty} e^{-\varsigma^\lambda} \varsigma^{\frac{n}{2}} J_{\frac{n}{2}-1}(2\pi \|\omega\| \varsigma) d\varsigma. \tag{10}$$

### 3.4. Fourier Transform and Maclaurin Series

If we perform the integration term by term in (9), we get

$$\hat{G}_\lambda(\omega) = 2\pi^{\frac{n}{2}} \sum_{k=0}^{+\infty} (-1)^k \frac{(\pi \|\omega\|)^{2k}}{\Gamma(k + 1) \Gamma(\frac{2k+n}{2})} \int_0^{+\infty} e^{-\varsigma^\lambda} \varsigma^{2k+n-1} d\varsigma.$$

Since

$$\int_0^{+\infty} e^{-\varsigma^\lambda} \varsigma^{2k+n-1} d\varsigma = \frac{1}{\lambda} \Gamma\left(\frac{2k+n}{\lambda}\right),$$

we finally have

$$\hat{G}_\lambda(\omega) = \frac{2\pi^{\frac{n}{2}}}{\lambda} \sum_{k=0}^{+\infty} (-1)^k \frac{\Gamma\left(\frac{2k+n}{\lambda}\right)}{\Gamma(k + 1) \Gamma\left(\frac{2k+n}{2}\right)} (\pi \|\omega\|)^{2k}. \tag{11}$$

The interchange of integral and summation symbols in these expressions is allowed as long as the series is uniformly convergent. To obtain the condition on  $\lambda$ , we use the ratio test and let

$$A_k = (-1)^k \frac{\Gamma\left(\frac{2k+n}{\lambda}\right)}{\Gamma(k + 1) \Gamma\left(\frac{2k+n}{2}\right)} \pi^{2k}.$$

Using the following asymptotic formula for the Gamma function (see [1])

$$\Gamma(az + b) = \sqrt{2\pi} e^{-az} (az)^{az+b-\frac{1}{2}} \left(1 + o\left(\frac{1}{z}\right)\right)$$

which holds for  $z \in \mathbb{C}$  such that  $|\arg(z)| < \pi$  and  $a > 0$ , and where

$$o\left(\frac{1}{z}\right) \rightarrow 0 \quad \text{for} \quad |z| \rightarrow +\infty,$$

we can write

$$\begin{aligned} \left| \frac{A_k}{A_{k+1}} \right| &= \frac{\Gamma\left(\frac{2k+n}{\lambda}\right)}{\Gamma\left(\frac{2k+n+2}{\lambda}\right)} \frac{(k+1)\left(k+\frac{n}{2}\right)}{\pi^2} \\ &= \left(\frac{\lambda}{2k+n}\right)^{\frac{2}{\lambda}} \frac{(2k)^2}{(2\pi)^2} \left(1+\frac{1}{k}\right)\left(1+\frac{n}{2k}\right)\left(1+o\left(\frac{1}{k}\right)\right) \\ &= \left(\frac{\lambda}{1+\frac{n}{2k}}\right)^{\frac{2}{\lambda}} \frac{(2k)^{2-\frac{2}{\lambda}}}{(2\pi)^2} \left(1+o\left(\frac{1}{k}\right)\right). \end{aligned}$$

Thus the radius of convergence is now obtained from

$$R_\lambda = \sqrt{\lim_{k \rightarrow +\infty} \left| \frac{A_k}{A_{k+1}} \right|} = \begin{cases} 0 & \text{for } 0 < \lambda < 1, \\ 1/2\pi & \text{for } \lambda = 1, \\ +\infty & \text{for } \lambda > 1. \end{cases}$$

It follows that the Maclaurin series (11) converges for  $\lambda > 1$  for any  $\omega \in \mathbb{R}^n$  and for  $\lambda = 1$  for any  $\omega \in \mathbb{R}^n$  such that  $\|\omega\| < 1/2\pi$ .

**Example 1.** For  $\lambda = 2$  we have

$$\hat{G}_2(\omega) = \pi^{\frac{n}{2}} \sum_{k=0}^{+\infty} (-1)^k \frac{(\pi \|\omega\|)^{2k}}{k!} = \pi^{\frac{n}{2}} e^{-(\pi \|\omega\|)^2}.$$

**Example 2.** For  $\lambda = 1$  and  $n = 1$  we have

$$\begin{aligned} \hat{G}_1(\omega) &= 2\pi^{\frac{1}{2}} \sum_{k=0}^{+\infty} (-1)^k \frac{\Gamma(2k+1)}{\Gamma(k+1)\Gamma\left(\frac{2k+1}{2}\right)} (\pi\omega)^{2k} \\ &= 2 \sum_{k=0}^{+\infty} (-1)^k (2\pi\omega)^{2k} \\ &= \frac{2}{1+(2\pi\omega)^2}, \end{aligned}$$

where the last expression is valid for any  $\omega \in \mathbb{R}$ .

**4. Other Extensions**

We consider the following more general form of (6) given by

$$GQ_\lambda(x) = e^{-(x^t Q x)^{\frac{\lambda}{2}}}, \tag{12}$$

where  $Q$  is a real positive-definite symmetric matrix of order  $n$ . Then its Fourier transform is given by

$$\widehat{GQ}_\lambda(\omega) = \int e^{-(x^t Q x)^{\frac{\lambda}{2}}} e^{-2\pi i \omega^t x} dx. \tag{13}$$

We can write  $Q = P^t D P$ , where  $P$  is an orthogonal matrix and  $D = \text{diag}(d_1, d_2, \dots, d_n)$  is a diagonal matrix with strictly positive real entries  $d_i > 0$ . We have

$$Q = P^t D P = P^t D^{\frac{1}{2}} D^{\frac{1}{2}} P, \tag{14}$$

where  $D^{\frac{1}{2}} = \text{diag}(d_1^{\frac{1}{2}}, d_2^{\frac{1}{2}}, \dots, d_n^{\frac{1}{2}})$ , then

$$x^t Q x = x^t P^t D^{\frac{1}{2}} D^{\frac{1}{2}} P x. \tag{15}$$

Let us set  $y = D^{\frac{1}{2}} P x$  or  $x = P^t D^{-\frac{1}{2}} y$ , then

$$\omega^t x = \omega^t P^t D^{-\frac{1}{2}} y = (D^{-\frac{1}{2}} P \omega)^t y, \tag{16}$$

and (13) becomes

$$\widehat{GQ}_\lambda(\omega) = \frac{1}{|\text{Det}(D^{\frac{1}{2}} P)|} \hat{G}_\lambda(D^{-\frac{1}{2}} P \omega), \tag{17}$$

since  $dy = |\text{Det}(D^{\frac{1}{2}} P)| dx$  and where  $\text{Det}(D^{\frac{1}{2}} P) = [\prod_{k=1}^n d_k]^{\frac{1}{2}} \text{Det}(P)$ .

Finally we could also use the  $p$ -norm  $\|x\|_p$  in  $\mathbb{R}^n$  given by

$$\|x\|_p = \left( \sum_{i=1}^n |x_i|^p \right)^{\frac{1}{p}}$$

and define another generalized Gaussian distribution by the formula

$$G_{\lambda,p}(x) = e^{-\|x\|_p^\lambda}.$$

Formulas like (10) and (11) are not yet established for this form of GG.



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### References

- [1] M. Abramowitz, I.A. Segun, *Handbook of Mathematical Functions*, Dover Publications, Inc., New York (1965).
- [2] G. Agrò, Maximum likelihood estimation for the exponential power function parameters, *Computational Statistics and Data Analysis*, **24** (1995), 523-536.
- [3] S. Basu, C.A. Micchelli, P. Olsen, Power exponential densities for training and classification of acoustics feature vectors in speech recognition, *Journal of Computational Graphical Statistics*, **10** (2001), 158-184.
- [4] G.E.P. Box, A note on regions for test of kurtosis, *Biometrika*, **40** (1953), 465-468.
- [5] G.E.P. Box, G. Tiao, A further look at robustness via Bayes's Theorem, *Biometrika*, **49** (1962), 419-432.
- [6] G.E.P. Box, G. Tiao, *Bayesian Inference in Statistical Analysis*, Reading, MA, Addison-Wesley (1973).
- [7] S.T.B. Choi, S.G. Walker, The extended exponential power distribution and Bayesian robustness, *Statistics and Probability Letters*, **65** (2003), 227-232.
- [8] A. Desgagné, J.-F. Angers, Importance sampling with the generalized exponential power density, *Statistics and Computing*, **15** (2005), 189-196.
- [9] P.H. Dianada, Note on some properties of maximum likelihood estimates, *Proceedings of the Cambridge Philosophical Society*, **45** (1949), 536-544.
- [10] T.J. DiCiccio, A.C. Monti, Inferential aspects of the skew exponential power distribution, *Journal of the American Statistical Association*, **99** (2004), 439-450.

- [11] J.A. Domínguez-Molina, G. González-Farías, R.M. Rodríguez-Dagnino, A practical procedure to estimate the shape parameter in the generalized Gaussian distribution, *Technical Report*, I-01-18, Investigation Center in Mathematics (CIMAT), Guanajuato, Mexico (2001).
- [12] F. Dubeau, S. El Mashoubi, Fourier transform of generalized Gaussian distribution for integer valued parameter, *Advances and Applications in Mathematical Sciences*, **1** (2009), 311-321.
- [13] F. Dubeau, S. El Mashoubi, Fourier transform of generalized Gaussian distribution for real valued parameter, *Advances and Applications in Mathematical Sciences*, **3** (2011), To Appear.
- [14] M. Fréchet, Sur la loi des erreurs d'observations, *Mathematischeskii Sbornik*, **32** (1924), 1-8.
- [15] E. Gómez, M.A. Gómez-Villegas, J.M. Marín, A multivariate generalization of the power exponential family of distributions, *Communications in Statistics, Theory and Methods*, **27** (1998), 589-600.
- [16] A. Hazan, Z. Landsman, U.E. Makov, Robustness via a mixture of exponential power distributions, *Computational Statistics and Data Analysis*, **42** (2003), 111-121.
- [17] H. Jakuszenkow, Estimation of the variance in the generalized Laplace distribution with quadratic loss function, *Demonstratio Mathematica*, **12** (1979), 581-591.
- [18] N.L. Johnson, S. Kotz, N. Balakrishnan, *Continuous Univariate Distributions*, Volume II, Wiley (1995).
- [19] D.W. Kammler, *A First Course in Fourier Analysis*, Prentice Hall, New Jersey (2000).
- [20] C. Kleiber, S. Kotz, *Statistical Size Distributions in Economics and Actuarial Sciences*, Wiley (2003).
- [21] T.W. Körner, *Fourier Analysis*, Cambridge University Press, Cambridge (1988).
- [22] S. Kotz, T.J. Kozubowski, K. Podgorski, *The Laplace Distribution and Generalizations*, Birkhäuser (2001).

- [23] G. Lunetta, Di una generalizzazione dello schema della curva normale, *Annali della Facolta di Economia e Commercio di Palermo*, **17** (1963), 237-244.
- [24] S. Metari, F. Deschênes, A new convolution kernel for atmospheric point spread function applied to computer vision, In: *IEEE 11-th Conference on Computer Vision (ICCV 2007)*, Rio de Janeiro, Brazil, 14-20 October (2007).
- [25] A.C. Monti, A note on estimation of the skew normal and skew exponential power distributions, *Metron*, **61** (2003), 205-219.
- [26] G. Moser, J. Zerubia, S.B. Serpico, SAR amplitude probability density function estimation based on a generalized Gaussian scattering model, *Technical Report*, 5153, INRIA (2004).
- [27] S. Nadarajah, A generalized normal distribution, *Journal of Applied Statistics*, **32** (2005), 685-694.
- [28] S. Nadarajah, Acknowledgement of priority: The generalized normal distribution, *Journal of Applied Statistics*, **33** (2006), 1031-1032.
- [29] E.G. Sánchez-Manzano, M.A. Gómez-Villegas, J.-M. Marín-Diazaraque, A matrix variate generalization of the power exponential family of distributions, *Communications in Statistics, Theory and Methods*, **31** (2002), 2167-2182.
- [30] D. Sharma, On estimating the variance of a generalized Laplace distribution, *Metrika*, **31** (1984), 85-88.
- [31] M.T. Subbotin, On the law of frequency of error, *Mathematicheskii Sbornik*, **31** (1923), 296-300.
- [32] P.R. Tadikamalla, Random sampling from the exponential power distribution, *Journal of the American Statistical Association*, **75** (1980), 683-686.
- [33] J.M.G. Taylor, Properties of modelling the error distribution with an extra shape parameter, *Computational Statistics and Data Analysis*, **13** (1992), 33-46.
- [34] G.C. Tiao, D.R. Lund, The use of OLUMV estimators in inference robustness studies of the location parameter of a class of symmetric distributions, *Journal of the American Statistical Association*, **65** (1970), 370-386.

- [35] C.A. Tojeiro, F. Louzada-Neto, H. Bolfarine, A Bayesian analysis for accelerated lifetime tests under an exponential power law model with threshold stress, *Journal of Applied Statistics*, **31** (2004), 685-691.
- [36] M.C. Turner, On heuristic estimation methods, *Biometrics*, **16** (1960), 299-301.
- [37] M.K. Varanasi, B. Aazhany, Parametric generalized Gaussian density function, *Journal of the Acoustical Society of America*, **86** (1989), 1404-1415.
- [38] S. Vianelli, La misura della variabilita condizionata in uno schema generale delle curve normali di frequenza, *Statistica*, **42** (1963), 155-176.
- [39] R. Zeckhauser, M. Thompson, Linear regression with non-normal error terms, *Review of Economics and Statistics*, **52** (1970), 280-286.