

**A NON-UNIFORM BOUND ON POISSON APPROXIMATION
FOR INDEPENDENT BERNOULLI RANDOM SUMMANDS**

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Abstract: A non-uniform bound for the distance between the distribution of random sums of independent Bernoulli random variables and an appropriate Poisson distribution is obtained. It is sharper than the bound reported in [7].

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1. Introduction

Let X_1, X_2, \dots be a sequence of independent Bernoulli random variables, each with probability $p_i = P(X_i = 1) = 1 - P(X_i = 0)$, and let N be a non-negative integer-valued random variable and independent of the X_i 's. Let us consider the random sums $S_N = \sum_{i=1}^N X_i$ and $N = n \in \mathbb{N}$ is fixed, there has been some research on topics related to the approximation of the distribution of S_n by an appropriate Poisson distribution, which can be found in [1], [2], [3]–[6], etc. Let $\lambda_N = \sum_{i=1}^N p_i$, $\lambda = E(\lambda_N)$ and U_λ a Poisson random variable with mean λ . For approximating the distribution of S_N , Yannaros [7] gave a uniform bound for the distance between the distribution of S_N and U_λ as follows:

$$|P(S_N \in A) - P(U_\lambda \in A)| \leq E|\lambda_N - \lambda| + E \left(\frac{1 - e^{-\lambda_N}}{\lambda_N} \sum_{i=1}^N p_i^2 \right), \quad (1.1)$$

where $A \subseteq \mathbb{N} \cup \{0\}$. It is observed that this bound is uniformly in A , which does not depend on A .

In this study, we are interest to determine a non-uniform bound of (1.1), which is in Section 2. Concluding remarks are presented in the last section.

2. Results

The following theorem presents a non-uniform bound for the distance between the distributions of S_N and U_λ , which is the desired result.

Theorem 2.1. *For $A \subseteq \mathbb{N} \cup \{0\}$, then we have*

$$|P(S_N \in A) - P(U_\lambda \in A)| \leq \min \left\{ 1, \sqrt{\frac{2}{e\lambda}} \right\} E|\lambda_N - \lambda| + \min \left\{ E \left(\frac{1 - e^{-\lambda_N}}{\lambda_N} \sum_{i=1}^N p_i^2 \right), \frac{1}{x_A} E \left(\sum_{i=1}^N p_i^2 \right) \right\}, \quad (2.1)$$

where $\frac{1}{x_A}$ is taken to be 1 when $x_A = 0$ ($x_A^\star = 0$ or $x_A^\star = 1$) and for $x_A > 0$, it is given by

$$\frac{1}{x_A} = \begin{cases} \frac{1}{x_A^\star} & \text{if } 0 \in A, \\ \frac{1}{x_A^\star - 1} & \text{if } 0 \notin A, \end{cases}$$

and $x_A^\star = \min\{x | x \in A\}$ and $x_A^\star = \max\{x | \{0, \dots, x\} \subseteq A\}$.

Proof. It follows the fact that

$$\begin{aligned} |P(S_N \in A) - P(U_\lambda \in A)| &\leq \sum_{n=0}^{\infty} P(N = n) |P(S_n \in A) - P(U_\lambda \in A)| \\ &\leq \sum_{n=0}^{\infty} P(N = n) |P(S_n \in A) - P(U_{\lambda_n} \in A)| \\ &\quad + |P(U_{\lambda_N} \in A) - P(U_\lambda \in A)|. \end{aligned} \quad (2.2)$$

Applying the corresponding results in [4] and [1], we obtain

$$|P(S_n \in A) - P(U_{\lambda_n} \in A)| \leq \min \left(\frac{1 - e^{-\lambda_n}}{\lambda_n} \sum_{i=1}^n p_i^2, \frac{1}{x_A} \sum_{i=1}^n p_i^2 \right)$$

and

$$|P(U_{\lambda_N} \in A) - P(U_\lambda \in A)| \leq \min \left\{ 1, \sqrt{\frac{2}{e\lambda}} \right\} E|\lambda_N - \lambda|,$$

respectively. Taking these bounds into (2.2), hence (2.1) is obtained. \square

If X_i 's are identically distributed, then the following corollary is an immediate consequence of the Theorem 2.1

Corollary 2.1. *Let $A \in \mathbb{N} \cup \{0\}$, if $p_1 = p_2 = \dots = p$, then we have the following:*

$$\begin{aligned} |P(S_N \in A) - P(U_\lambda \in A)| &\leq \min \left\{ E(1 - e^{-pN})p, \frac{p^2 E(N)}{x_A} \right\} \\ &\quad + \min \left\{ 1, \sqrt{\frac{2}{eE(N)p}} \right\} pE|N - E(N)|. \end{aligned} \quad (2.3)$$

Remark. We consider the following facts:

1. $\min \left(\frac{1 - e^{-\lambda_N}}{\lambda_N}, \frac{1}{x_A} \right) \leq \frac{1 - e^{-\lambda_N}}{\lambda_N}$ and
2. $\min \left\{ 1, \sqrt{\frac{2}{e\lambda}} \right\} E|\lambda_N - \lambda| \leq E|\lambda_N - \lambda|.$

Thus, the bound in (2.1) is sharper than the bound in (1.1).

3. Conclusion

In this study, a non-uniform bound for the distance between the distribution of random sums of independent Bernoulli random variables and an appropriate Poisson distribution was determined. By comparing, this bound is sharper than that reported in [7]. Thus, it is more appropriate for measuring the accuracy of the approximation.

References

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