

**A NOTE ON THE ZUBAIR–G FAMILY WITH BASELINE  
LOMAX CUMULATIVE DISTRIBUTION FUNCTION.  
SOME APPLICATIONS**

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**Abstract:** In this paper we study the one–sided Hausdorff approximation of the shifted Heaviside step function by a class of the Zubair–G family of cumulative lifetime distribution with baseline Lomax c.d.f. The estimates of the value of the best Hausdorff approximation obtained in this article can be used in practice as one possible additional criterion in "saturation" study.

As an illustrative example we consider the modelling of the growth of red abalone (*Haliotis Rufescens*) in Northern California.

We also look at a possible extension, which we call  $\alpha$ –Zubair–G Family with baseline Lomax (cdf).

Finally, the potentiality of the new software reliability model analyzed by means of real dataset. Some comparisons are made. Numerical examples, illustrating our results are presented using programming environment *CAS Mathematica*.

**AMS Subject Classification:** 68N30, 41A46

**Key Words:** Zubair–G family of cumulative lifetime distribution, Zubair–G Family with baseline Lomax (cdf),  $\alpha$ –Zubair–G Family with baseline Lomax (cdf), Heaviside function, Hausdorff approximation, upper and lower bounds

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## 1. Introduction

In [1], a new family of lifetime distributions, called the Zubair–G family of distributions is introduced.

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The new family is defined by the following cumulative distribution function (cdf)

$$F(t; \lambda) = \frac{e^{\lambda G^2(t)} - 1}{e^\lambda - 1}, \quad (1)$$

where  $\lambda > 0$ .

If,  $G$  is the (cdf) of the baseline model, then the distribution function (1) will be the (cdf) of the Zubair-G family.

Some comments on a Zubair-G Family of cumulative lifetime distributions with baseline Weibull (cdf) can be found in [2].

We also look at a possible extension, which we call  $\alpha$ -Zubair-G Family.

For example, if  $G(t)$  be (cdf) of the Lomax distribution [10] given by

$$G(t) = 1 - \frac{1}{1 + \left(\frac{t}{b}\right)^a} \quad (2)$$

then the (cdf) of the Z-Lomax distribution has the form

$$F(t) = \frac{e^{\lambda \left(1 - \frac{1}{1 + \left(\frac{t}{b}\right)^a}\right)^2} - 1}{e^\lambda - 1} \quad (3)$$

where  $a > 0$  and  $b > 0$ .

We consider the following class of this family with application to the population dynamics and debugging theory:

$$M(t) = \frac{e^{\lambda \left(1 - \frac{1}{1 + \left(\frac{t}{b}\right)^a}\right)^2} - 1}{e^\lambda - 1}, \quad (4)$$

with

$$t_0 = b \left( \frac{\sqrt{\frac{1}{\lambda} \ln \frac{e^\lambda + 1}{2}}}{1 - \sqrt{\frac{1}{\lambda} \ln \frac{e^\lambda + 1}{2}}} \right)^{\frac{1}{a}}; \quad M(t_0) = \frac{1}{2}. \quad (5)$$

Some extensions of the well-known Poisson, Poisson-exponential, Chen, Exponentiated Chen, modified Weibull and Burr distributions can be found in: [3]-[14].

For some approximation, computational and modelling aspects, see [15]–[31].

Some software reliability models, can be found in [32]–[51], [58]–[59].

In this note we study the Hausdorff approximation of the *shifted Heaviside step function*

$$h_{t_0}(t) = \begin{cases} 0, & \text{if } t < t_0, \\ [0, 1], & \text{if } t = t_0, \\ 1, & \text{if } t > t_0 \end{cases}$$

by this family.

**Definition 1.** [52] The Hausdorff distance (the H-distance)  $\rho(f, g)$  between two interval functions  $f, g$  on  $\Omega \subseteq \mathbb{R}$ , is the distance between their completed graphs  $F(f)$  and  $F(g)$  considered as closed subsets of  $\Omega \times \mathbb{R}$ .

More precisely,

$$\rho(f, g) = \max\left\{ \sup_{A \in F(f)} \inf_{B \in F(g)} \|A - B\|, \sup_{B \in F(g)} \inf_{A \in F(f)} \|A - B\| \right\},$$

wherein  $\|\cdot\|$  is any norm in  $\mathbb{R}^2$ , e. g. the maximum norm  $\|(t, x)\| = \max\{|t|, |x|\}$ ; hence the distance between the points  $A = (t_A, x_A)$ ,  $B = (t_B, x_B)$  in  $\mathbb{R}^2$  is  $\|A - B\| = \max(|t_A - t_B|, |x_A - x_B|)$ .

We recall that completed graph of  $f$  is the closure of the graph of  $f$  as a subset of  $\Omega \times \mathbb{R}$ . If the graph of an interval function  $f$  equals  $\mathcal{F}(f)$ , then the  $f$  is called S-continuous.

The Hausdorff distance  $\rho(f, g) = \max\{\overrightarrow{\rho}(f, g), \overleftarrow{\rho}(g, f)\}$  defines a metric in the set of the S-continuous interval functions [53]–[56].

We propose a software modules (intellectual properties) within the programming environment CAS Mathematica for the analysis.

The models have been tested with real data (the growth of red abalone (*Haliotis Rufescens*) in Northern California).

Finally, the potentiality of the new software reliability model analyzed by means of real dataset.

Some comparisons are made.

### 2. Main Results

The one-sided Hausdorff distance  $d$  between the function  $h_{t_0}(t)$  and the sigmoidal function - ((4)–(5)) satisfies the relation

$$M(t_0 + d) = 1 - d. \tag{6}$$

The following theorem gives upper and lower bounds for  $d$

**Theorem 1.** Let

$$p = -\frac{1}{2},$$

$$q = 1 + \frac{a\lambda(1+e^\lambda)\sqrt{\frac{1}{\lambda} \ln \frac{e^\lambda+1}{2}} \left(1 - \sqrt{\frac{1}{\lambda} \ln \frac{e^\lambda+1}{2}}\right)^2}{b(e^\lambda-1)} \times$$

$$\left(\frac{\sqrt{\frac{1}{\lambda} \ln \frac{e^\lambda+1}{2}}}{1 - \sqrt{\frac{1}{\lambda} \ln \frac{e^\lambda+1}{2}}}\right)^{\frac{a-1}{a}}, \tag{7}$$

$$r = 2.1q.$$

For the one-sided Hausdorff distance  $d$  between  $h_{t_0}(t)$  and the sigmoid ((4)–(5)) the following inequalities hold for:  $q > \frac{e^{1.05}}{2.1}$

$$d_l = \frac{1}{r} < d < \frac{\ln r}{r} = d_r. \tag{8}$$

**Proof.** Let us examine the function:

$$F(d) = M(t_0 + d) - 1 + d. \tag{9}$$

From  $F'(d) > 0$  we conclude that function  $F$  is increasing.

Consider the function

$$G(d) = p + qd. \tag{10}$$

From Taylor expansion we obtain  $G(d) - F(d) = O(d^2)$ .

Hence  $G(d)$  approximates  $F(d)$  with  $d \rightarrow 0$  as  $O(d^2)$  (see Fig. 1).

In addition  $G'(d) > 0$ .

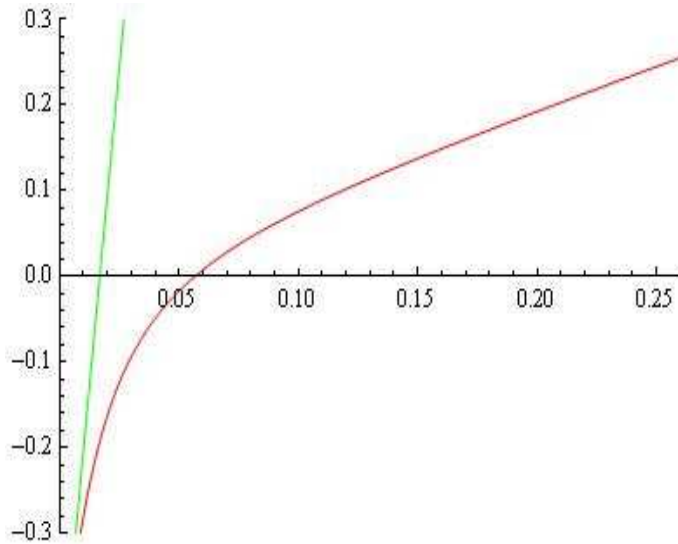


Figure 1: The functions  $F(d)$  and  $G(d)$  for  $b = 0.01$ ;  $a = 1.95$   $\lambda = 0.95$ .

Further, for  $q > \frac{e^{1.05}}{2.1}$  we have  $G(d_l) < 0$  and  $G(d_r) > 0$ .

This completes the proof of the theorem.

### 3. Numerical examples.

The model ((4)–(5)) for  $b = 0.05$ ;  $a = 1.9$ ;  $\lambda = 0.9$ ,  $t_0 = 0.0974305$  is visualized on Fig. 2.

From the nonlinear equation (6) and inequalities (8) we have:  $d = 0.139818$ ,  $d_l = 0.0725397$ ,  $d_r = 0.190317$ .

The model ((4)–(5)) for  $b = 0.01$ ;  $a = 1.95$   $\lambda = 0.95$ ,  $t_0 = 0.0193646$  is visualized on Fig. 3.

From the nonlinear equation (6) and inequalities (8) we have:  $d = 0.0566589$ ,  $d_l = 0.01602$ ,  $d_r = 0.0662253$ .

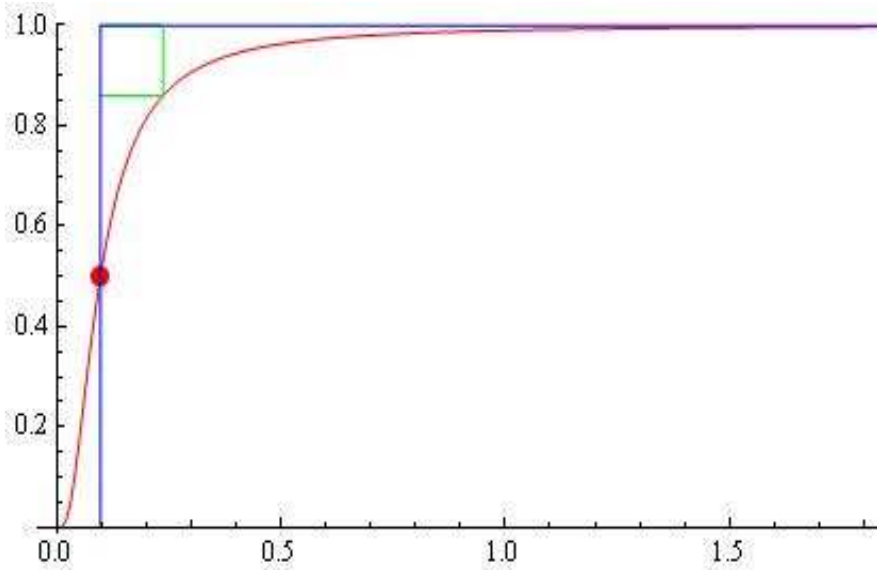


Figure 2: The model ((4)–(5)) for  $b = 0.05$ ;  $a = 1.9$ ;  $\lambda = 0.9$ ,  $t_0 = 0.0974305$ ; H-distance  $d = 0.139818$ ,  $d_l = 0.0725397$ ,  $d_r = 0.190317$ .

From the above examples, it can be seen that the proven estimates (see Theorem 1) for the value of the Hausdorff approximation is reliable when assessing the important characteristic - "saturation".

This characteristic (as we have already shown in our previous publications) has its equal participation together with the other two characteristics - "confidence intervals" and "confidence bounds" in the area of the Software Reliability Theory.

We propose a software module (intellectual properties) within the programming environment *CAS Mathematica* for the analysis of the considered family  $M(t)$ .

The module offers the following possibilities:

- generation of the function under user defined values of the parameters  $a$ ,  $\lambda$ , and  $b$ ;
- calculation of the H-distance  $d$  between the function  $h_{t_0}(t)$  and the function  $M(t)$ ;
- software tools for animation and visualization.

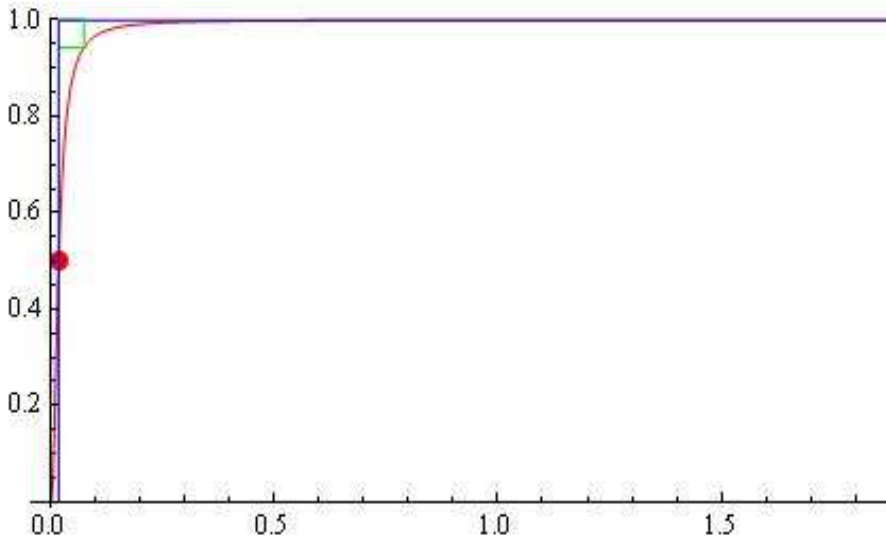


Figure 3: The model ((4)–(5)) for  $b = 0.01$ ;  $a = 1.95$   $\lambda = 0.95$ ,  $t_0 = 0.0193646$ ; H-distance  $d = 0.0566589$ ,  $d_l = 0.01602$ ,  $d_r = 0.0662253$ .

#### 4. Some applications.

##### 4.1. Application to the Population Dynamics

We examine the following data. (The data for modeling the growth of red abalone is shown in Table 1. For more details, see [57]).

The model

$$M(t) = \omega \left( \frac{e^{\lambda \left(1 - \frac{1}{1 + (\frac{t}{b})^a}\right)^2} - 1}{e^\lambda - 1} \right)$$

based on the data of Table 1 for the estimated parameters:

$$\omega = 242.5; \quad b = 1.7205; \quad \lambda = 2.17481; \quad a = 1.39039$$

is plotted on Fig. 4.

<i>Age</i>	<i>Length(mm)</i>
1	16.1
2	33.9
3	54.3
4	76.2
5	97.8
6	117.1
7	133.3
8	146.5
9	157.2
10	166
11	173.3
12	179.6
13	185
14	189.5
15	194

Table 1: Data for modeling the growth of red abalone *Haliotis Rufescens* in Northern California [57]

For the predictive power (PP) criterion:

$$PP = \sum_{i=1}^n \left( \frac{M(t_i) - y_i}{y_i} \right)^2$$

measures the distance of model actual data from the estimates against the actual data, we find  $PP = 0.29094$ .

We hope that the results will be useful for specialists in this scientific area.

#### 4.2. Application to the Software Reliability Growth Theory

We examine the following data.

The fitted model  $M(t)$  based on the data of Table 2 for the estimated parameters:

$$\omega = 6543; \quad b = 4.07363; \quad \lambda = 2.33458; \quad a = 1.16652$$



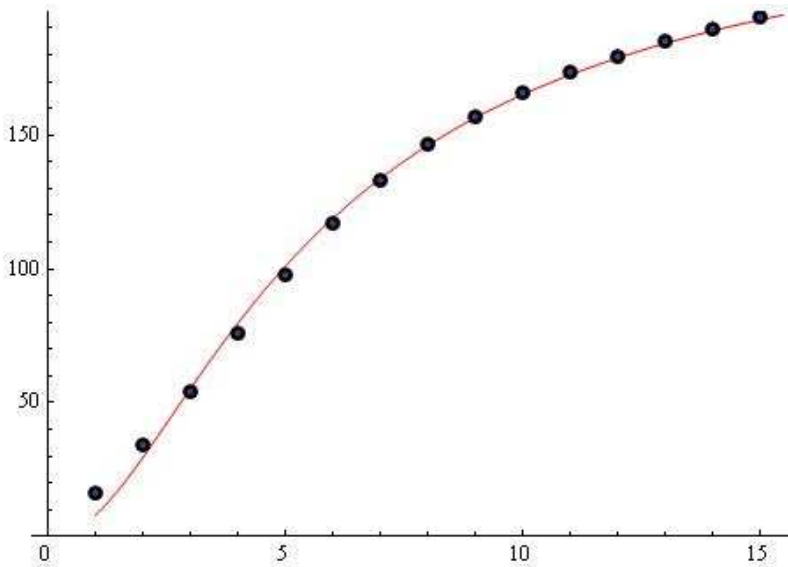


Figure 4: The model  $M(t)$  based on the data of Table 1

is plotted on Fig. 5.

We will explicitly note that in some cases the presented software reliability model provides better results than other much more sophisticated models.

Comparison between the models  $M(t)$  and the Gompertz software reliability model show a good fit by the presented new model (see, Fig. 6).

### 5. Concluding remarks.

Following the ideas given in [2] we consider the *new  $\alpha$ -Zubair-G Family* with baseline Lomax (cdf):

$$M_{\alpha}(t) = \omega \left( \frac{e^{\lambda \left( 1 - \frac{1}{1 + \left(\frac{t}{b}\right)^{\alpha}} \right)^2} - 1}{e^{\lambda} - 1} \right)^{\alpha}, \tag{11}$$

where  $\alpha > 0$ .

<i>Week</i>	<i>Cumulative number of software faults</i>
1	248
5	742
10	1936
15	2567
20	3205
25	3750
30	4301
35	4548
40	4749
45	4901
50	5024
55	5110

Table 2: The actual data (week; cumulative number of software faults) [61], [60]

The model  $M_\alpha(t)$  based on the data of Table 1 for the estimated parameters:  $\omega = 194$ ;  $b = 7.88871$ ;  $\lambda = 2.2408$ ;  $a = 4.71474$ ;  $\alpha = 0.110994$  is plotted on Fig. 7.

For the predictive power (PP) criterion:

$$PP = \sum_{i=1}^n \left( \frac{M_\alpha(t_i) - y_i}{y_i} \right)^2$$

we find  $PP = 0.0035919$ .

Comparison between the models  $M(t)$  and  $M_\alpha(t)$  show a good fit by the presented new model  $M_\alpha(t)$  (see, Fig. 7).

Based on the methodology proposed in the present note, the reader may formulate the corresponding approximation problems for the general model  $M_\alpha(t)$  (11) on his/her own.

```

Manipulate[Dynamic@Show[Plot[f[t], {t, 1, 55}, LabelStyle -> Directive[Green, Bold],
  PlotLabel -> 6543 * (Exp[λ * (1 - 1 / (1 + (t / b)^a))^2] - 1) / (Exp[λ] - 1)],
  ListPlot[{{1, 248}, {5, 742}, {10, 1936}, {15, 2567}, {20, 3205}, {25, 3750},
    {30, 4301}, {35, 4548}, {40, 4749}, {45, 4901}, {50, 5024}, {55, 5110}},
  PlotStyle -> Red, PlotMarkers -> {Automatic, Small}],
  PlotRange -> {Automatic, {0, 5100}}, AxesOrigin -> {0, 0}],
{{b, 0}, 0.1, 10, Appearance -> "Open"}, {{a, 0}, 0.01, 10, Appearance -> "Open"},
{{λ, 0}, 0.01, 10, Appearance -> "Open"},
Initialization -> {f[t_] := 6543 * (Exp[λ * (1 - 1 / (1 + (t / b)^a))^2] - 1) / (Exp[λ] - 1)}]

```

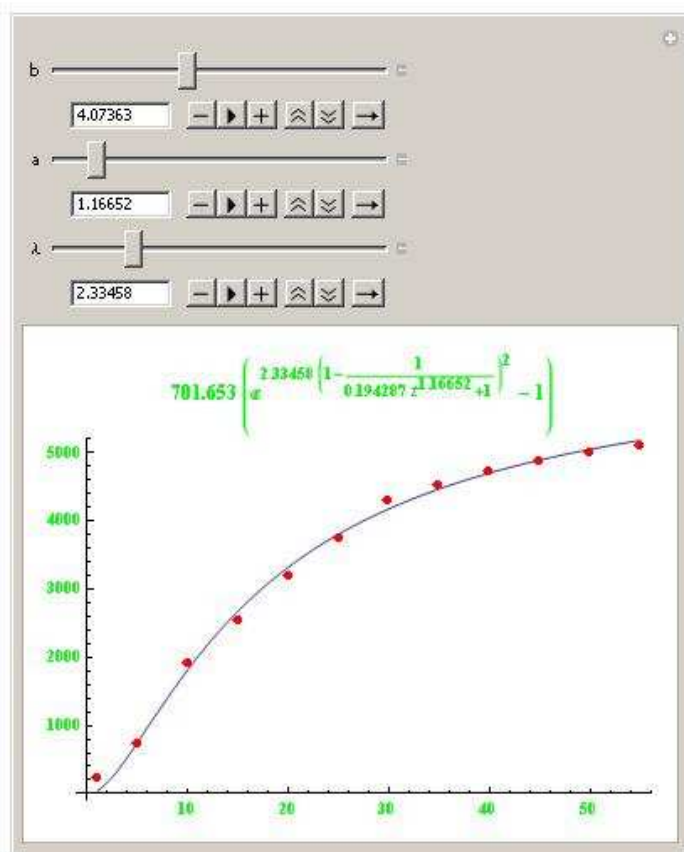


Figure 5: The model  $M(t)$  based on the data of Table 2.

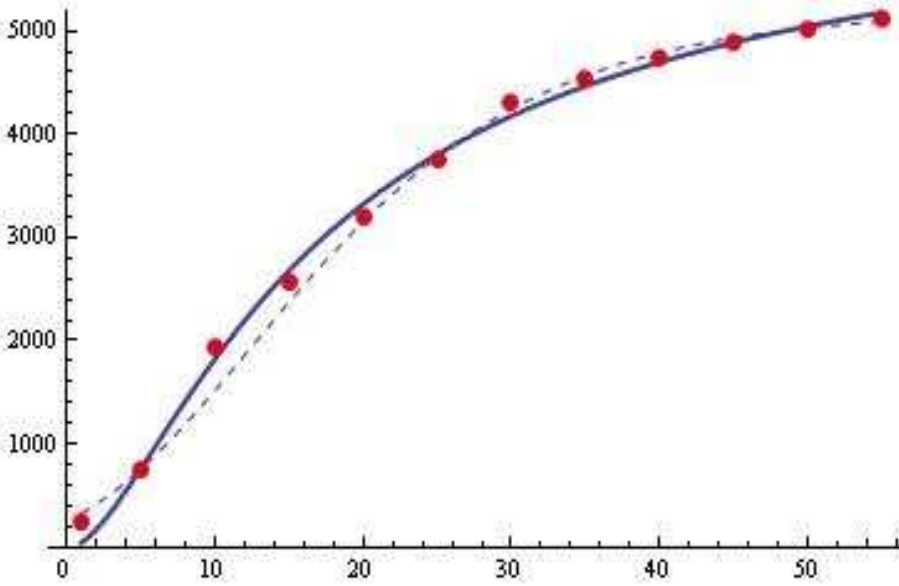


Figure 6: The models:  $M(t)$  (thick) and Gompertz software reliability model (dashed) based on the data of Table 2.

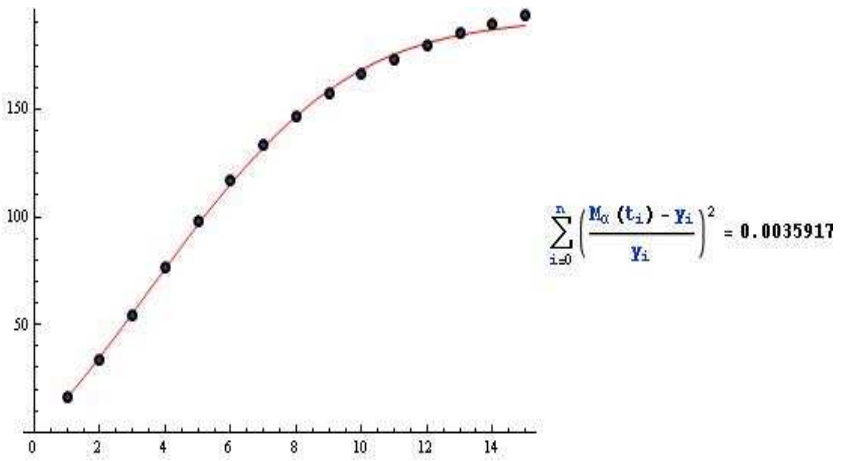


Figure 7: The model  $M_{\alpha}(t)$  based on the data of Table 1.

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